ABSTRACT

Counting abilities have been described as determinative precursors for a good development of later mathematical abilities. However, an important part of variance in mathematical achievement has also been associated with differences between instruction given in schools.

In this study counting and instruction as predictors for mathematical skills were studied in 423 children. Our data revealed that the mastery of the counting principles in kindergarten was predictive for the risk for mathematical (dis)abilities in grade 1. Moreover, children sharing a common instructional background tended to have more similar scores on mathematical tests, yet the importance of mastery of the counting principles in the prediction of later mathematical achievement was the same for all classrooms.

Key words: mathematical abilities, risk for math disability, counting, procedural calculation, fact retrieval, kindergarten predictors, classroom, instruction
Mathematics instruction: do classrooms matter?

**INTRODUCTION**

Mathematics is inherently present in everyday life; each day we are confronted with it such as paying in the shop, baking a cake, travelling by train... Although mathematical problems have serious educational consequences, this area has received little attention in research until recently (Engle, Grantham-McGregor, Black, Walker, & Wachs, 2007; Landerl, Bevan, & Butterworth, 2004). However, Dowker (2005) indicated that the impact of poor mathematical skills is greater than the influence of poor reading skills.

Differences in mathematics between and within individuals are normal. Teachers are expected to cope with learning differences and to adjust their teaching style to the needs of all students. However in some cases these differences appear to be so severe or resistant that they can be considered as characteristics of ‘problems’ or even ‘disabilities’ (Grégoire & Desoete, 2009). Most practitioners and researchers currently report a prevalence of Mathematical Learning Disabilities (MLD) between 2-14% of children (Barbaresi, Katuskic, Colligan, Weaver, & Jacobsen, 2005; Desoete, Roeyers, & De Clercq, 2004; Shalev, Manor, & Gross-Tsur, 2005). The prevalence of MLD in siblings even ranges from 40 to 64% (Desoete, Praet, & Ceulemans, 2013; Shalev et al., 2001).

The term MLD refers to a significant degree of impairment in the mathematical skills (with substantially below mathematical performances). In addition, children with MLD do not profit enough from (good) help. This is also referred to as a lack of Responsiveness to intervention. Finally, the problems in MLD can not be totally explained by impairments in general intelligence or external factors that could provide sufficient evidence for scholastic failure. Persons with MLD describe their problems as follows (Desoete, Van Hees, Tops, & Brysbaert, 2012; Vanmeirhaeghe & VanHees, 2012):
Kristel (master in education) : “Why was elementary school like hell? Because I felt a huge pressure on me. Open your manual on page 68. There we go again! Where is page 68? Other pupils already had taken down the title while I was still looking for page 68. It was a constant feeling of needing to exert myself. I have to take care that I can follow. That is what made it so hard for me. Everyone was faster then I was. “

A child with MLD needs extra support to keep following teaching according to its own intellectual level. MLD goes namely beyond (mental) arithmetic. Even remembering definitions takes more efforts.

Sara (bachelor in journalism) I need three times more time than an average student to learn the same subjects.

While early literacy is stimulated by almost all parents, early numeracy and counting gets less universal attention, although also the development of mathematical (dis)abilities begins before formal schooling starts (Ceulemans, Loeys, Warreyn, Hoppenbrouwers, & Desoete, 2012; Sophian, Wood, & Vong, 1995). It is therefore not surprising that children start with a quite heterogeneous baggage of counting skills at the school-desk. In addition, Opdenakker and Van Damme (2006) found that an important part of the variance in mathematical abilities in first grade were associated with differences between schools.

This study focused on counting abilities as predictor (Aunola et al., 2004; Gersten, Jordan, & Flojo, 2005; Le Fevre et al., 2006) in combination with mathematical instruction (Opdenakker & Van Damme, 2006) in the prediction of mathematical (dis)abilities in a large sample of children with a wide range of (dis)mathematical abilities.

Counting
Counting can be considered as a key ability for the development of age adequate mathematical skills. By means of counting, number facts are stored in long-term memory (Geary, 2011). In addition, counting activities lead to better strategies for addition and subtraction (Le Fevre et al., 2006) and multiplication (Blöte, Lieffering, & Ouwehand, 2006).

The mastery of the essential counting principles has been described as an essential feature for the development of counting (Geary, 2004; Gelman & Meck, 1983; Wynn, 1992). Children have to master the stable order, the one-one correspondence and the cardinality principle in kindergarten. The stable order principle implies that the order of number words must be invariant across counted sets. The one-one-correspondence principle holds that every number word can only be attributed to one counted object. Once the cardinality principle is acquired, children know that the value of the last number word represents the quantity of the counted objects. Knowledge of the stable-order principle is reliable first of all, followed by the one-one correspondence principle, while mastery of the cardinality principle was found to develop the slowest (Butterworth, 2004; Fuson, 1988).

**Mathematical instruction**

Mathematical instruction might differ in the adopted instructional paradigm (Case, 1998; Daniels & Shumow, 2003; De Corte, 2004; Ellis & Berry, 2005). The adoption of a traditional approach (e.g., emphasis on rules, memorizing and rehearsing), a structuralist approach (e.g., stressing abstract conceptualizations of mathematical content) or a constructivistic view towards learning (e.g., teaching mathematics presenting problems within a familiar context in order to give meaning), will affect the design of learning materials and the instructional strategies suggested in textbooks (Carnine, Dixon, & Silbert, 1998; Van de Walle, 2007). This has been researched in an extensive way in relation to mathematics (Cooper, 1993; Nathan, Long, & Alibali, 2002). Moreover, differences between the
instructional interventions and curricula are found in the timing and the stage at which the conceptions are presented to children as well as in the kinds of learning opportunities provided and in its organizing and sequencing (Schmidt, McKnight, Valverde, & Houang, 1997). As such, a large variation of teaching practices is adopted to teach mathematics in primary education. Depending on the curriculum, the textbooks used in the classroom, and the preferences and beliefs of each individual teacher, instruction can strongly differ across classrooms (Remillard, 1999). However Slavin and Lake (2008) revealed that there is a lack of evidence supporting a differential effect of mathematics curricula on students’ mathematics performance results.

**Objectives**

Although some authors stressed the importance of instruction and curricula (e.g., Chval, Chávez, Reys, & Tarr, 2009; Van Steenbrugge, Valcke, & Desoete, 2010; Zhao, Valcke, Desoete, Verhaeghe, & Xu, 2011) there is inconclusive evidence (Slavin & Lake, 2008) on the influence of instruction on children’s mathematical skills in grade 1.

In this study the relationship between mastery of the counting principles in kindergarten (child factors) on the one hand and instruction (classroom factors) on the other hand on (dis) mathematical abilities will be studied.

**Method**

**Participants**

This study was carried out on 423 children (223 girls) in kindergarten. Of this sample 369 children were tested in grade 1. All children were Caucasian native Dutch-speaking children living in the Flemish part of Belgium. The children in this study had a mean age of
70.02 months ($SD = 4.01$ months) and attended on average 7.42 months ($SD = 1.03$ months) of school in the last kindergarten class when tested the first time.

Children were retrospectively classified as at risk for math learning disability (MLD) if they had scored $< -1.5$ on the Z-score of one of the mathematical ability tests in grade 1 ($n = 48$). Children who scored z-scores above $-1.5$ on both mathematical tests in grade 1 were classified as typical achievers ($n = 321$) not at risk for math disability.

**Materials**

All counting abilities were tested in kindergarten with the TEDI-MATH (Grégoire, Noel, & Van Nieuwenhoven, 2004). The TEDI-MATH has proven to be a well validated and reliable instrument. Children had to judge the counting of linear and random patterns of drawings and counters. To assess the abstraction principle, children had to count different kind of objects who were presented in a heap. Furthermore, a child who counted a set of objects was asked ‘how many objects are there in total?’, or ‘how many objects are there if you start counting with the leftmost object in the array’. When children had to count again to answer they did not gain any points, as this was considered to represent good procedural knowledge but a lack of understanding of the counting principles. One point was given for a correct answer with a correct motivation. A sum score was constructed (maximum: 13 points). Cronbach’s alpha was .85.

In order to obtain a complete overview of the mathematical abilities of children and to test for procedural calculation and semantic memory abilities (Pieters et al., 2013), the following mathematical tests were used: the Arithmetic Number Fact Test (Tempo Test Rekenen [TTR]; De Vos, 1992) and the Kortrijk Arithmetic Test Revision (Kortrijkse Rekentest-Revisie [KRT-R]; Baudonck et al., 2006).
The Arithmetic Number Fact Test (Tempo Test Rekenen [TTR]; De Vos, 1992) is a test consisting of number fact problems (e.g., \(2 + 5 = \ldots\); \(9 - 2 = \ldots\)). Children have to solve as many additions and subtractions as possible within 2 minutes. The psychometric value of the test has been demonstrated on a sample of 10,059 children.

The Kortrijk Arithmetic Test Revision (Kortrijkse Rekentest-Revisie [KRT-R]; Baudonck et al., 2006) is an untimed standardized test on procedural calculation from grade 1 until 6. The KRT-R requires that children solve calculations in a number-problem format (e.g., \(16 - 12 = \ldots\)) or in a word-problem format (e.g., 1 more than 3 is \ldots). The psychometric value of the test has been demonstrated on a sample of 3,246 children and is frequently used in Flemish education and diagnostic assessment.

**Procedure**

The children were recruited in 25 randomly selected schools, 9 schools were located in a city while 16 of them were located rurally. All parents received a letter with the explanation of the research and could submit informed consent in order to participate.

Children were tested during school time in a separate and quiet room. Toddlers were tested individually. The test leaders all received training in the assessment and interpretation of the tests. After completion of the test procedure, all the parents of the children received individual feedback on the results of their children.

**RESULTS**

In this sample only 44.2 % of children mastered the three counting principles by the end of kindergarten (see also Stock, Desoete, & Roeyers, 2009).

In addition, the MANOVA with procedural and conceptual counting skills as dependent variable and group (children at risk for MLD, children not at risk for MLD) as
group was significant on the multivariate level \((F (2, 366) = 37.241; p < .001, \text{ partial } \eta^2 = .169)\). There were significant differences on the univariate level for procedural \((F (1, 367) = 49.288; p < .001, \text{ partial } \eta^2 = .118)\) and for conceptual counting \((F (1, 367) = 48.832; p < .001, \text{ partial } \eta^2 = .117)\) with children at risk for math disability having lower developed procedural \((M = 41.31; \text{ SD } = 22.89)\) and conceptual \((M = 34.27 \text{ ; SD } = 28.29)\) counting abilities compared to peers not at risk for math disabilities (procedural counting \(M = 69.25; \text{ SD } = 26.10\); conceptual counting \(M = 64.35 \text{ ; SD } = 27.74\)).

Since the children in this study were clustered in classrooms and thus not sampled randomly and independently, intraclass correlations were computed for both dependent mathematical ability variables (the procedural calculation and fact retrieval skills of children in grade 1). The intraclass correlation was calculated as the proportion of the between-group variance relative to the sum of the between- and within-group variance.

Table 1

*Mixed Model Analysis: Null Model of mathematical abilities*

<table>
<thead>
<tr>
<th>Parameter</th>
<th>Procedural calculation</th>
<th>Numerical Facility</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Estimate</td>
<td>SE</td>
</tr>
<tr>
<td>Level 1 Intercept</td>
<td>.56*</td>
<td>.04</td>
</tr>
<tr>
<td>Level 2 Intercept</td>
<td>.57*</td>
<td>.18</td>
</tr>
<tr>
<td>Intraclass correlation</td>
<td>.50</td>
<td>.40</td>
</tr>
</tbody>
</table>

*Note.* \(p < .001\)

The intraclass indices (see Table 1) indicated that between 40 and 50% of the variance in the mathematical abilities of children could be explained by getting the same instruction. The individual level intercept variance was .56 for procedural calculation and .63 for numerical
facility. The classroom level intercept variance was .57 for procedural calculation and .42 for numerical facility.

In order to take into account this data structure, multilevel analyses were performed with counting skills as the independent predictor, the scores on the mathematical tests as Level 1 and classrooms (or instruction) as Level 2. The results of the analyses in Table 2.

Table 2

Mixed Model Analyses: Model including counting and mathematical instruction as factors

<table>
<thead>
<tr>
<th>Parameter</th>
<th>Procedural calculation</th>
<th>Numerical Facility</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Estimate</td>
<td>SE</td>
</tr>
<tr>
<td><strong>Fixed</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Intercept</td>
<td>-.13</td>
<td>.13</td>
</tr>
<tr>
<td>Counting skills</td>
<td>.33*</td>
<td>.06</td>
</tr>
<tr>
<td><strong>Random</strong></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Level 2 Intercept</td>
<td>.41*</td>
<td>.13</td>
</tr>
<tr>
<td>Level 2 Instruction</td>
<td>.03</td>
<td>.02</td>
</tr>
</tbody>
</table>

* p < .001

The fixed part of the model revealed that counting skills play a significant role in the prediction of both procedural calculation and numerical facility. Children with better counting skills in kindergarten tended to perform better on arithmetic tests in first grade. The random part of the model revealed that here was significant intercept variance between the classrooms for both arithmetic tests, indicating that classrooms differ in their mean performances. Yet no significant slope variance was found for the scores on the arithmetic tests between the different classrooms.
DISCUSSION

The goal of the current research was to gain more insight into the importance of mastering the counting principles in kindergarten versus the variance between classrooms or the role of instruction on mathematical abilities and the risk for math disability in grade 1.

In this study more than half of children did not master the three counting principles by the end of kindergarten. Big differences in the mastery of the essential counting principles in toddlers existed, so teachers may need to pay a lot of attention to the different baggage children bring with them when entering first grade.

In addition counting abilities in toddlers and their procedural calculation and fact retrieval abilities one year later in first grade were assessed in a large sample that included children with a wide range of mathematical abilities. Our findings revealed that it was possible to differentiate children at risk and not at risk for mathematical disabilities in elementary school based on the procedural and conceptual knowledge of counting in kindergarten.

Furthermore, it was supposed that children who did it better on the items of the counting principles as a whole in kindergarten, had better scores on mathematical tests in first grade one year later than children who had lower scores on the counting items. Since high values were found for the intraclass correlations, it was necessary to take into account the clustered structure of the data and to use multilevel analyses. The expected hypothesis could be confirmed. The better children performed on the counting items in the last kindergarten class, the better they performed on the two mathematical tests in first grade. These results confirm the role of counting abilities in the development of proficient arithmetic strategies (Stock et al., 2009, Stock, Desoete, & Roeyers, 2007; Van De Rijt & Van Luit, 1999).

The results pointed out that a large part of the variance in mathematical achievement in first grade can be associated with differences between schools. By using multilevel analyses it was possible to allow for similarities in the performances of children in the same
classroom, but no explanatory factors could be found. We found significant random variation for the mean class achievement indicating that the level of performances was quite different between schools and that children sharing a common educational background tended to have more similar scores on mathematical tests when compared with children in other schools. Yet there was no random slope variation, meaning that the importance of mastery of the counting principles in the prediction of later arithmetic achievement was the same for all classrooms. There was no differential influence of the school context on the children’s basic counting knowledge.

Yet the study had a few limitations. In this study the TEDI-MATH items (Grégoire et al., 2004) were used. We thus still have to be careful with our conclusions since MLD might not be a homogeneous disability (Pieters et al, 2013) and the choice of the used task can have an important impact on the results. Furthermore, the conclusions of this study have to be interpreted carefully since a large proportion of the variance remained unexplained. A lot of other possible powerful predictors besides the counting abilities such as language (e.g., Praet, Titeca, Ceulemans, & Desoete, 2013; Vukovic & Lesaux, 2013) and magnitude estimation skills (e.g., Stock, Desoete, & Roeyers, 2010) were not taken into account in this research. For example context variables such as home environment, and parental involvement (e.g., Reusser, 2000) should be included in future studies. These limitations indicate that only a part of the picture is investigated so the results of the study have to be interpreted with care. Yet the large group of children that was assessed in this study strengthens the generalizability of the results.

In conclusion our results revealed a relationship between mastery of the counting principles in kindergarten (child factors) on the one hand and instruction (classroom factors) on the other hand on mathematical abilities and the risk for math disability. It was possible to explain significant proportions of scores on mathematical tests in first grade based on the counting scores in kindergarten. In addition, there were important differences between
schools. Taking into account the large differences in baggage in terms of counting skills children took with them when starting basic schooling and the fact that scores on counting tasks were good predictors for later arithmetic abilities, it is important that teachers in first grade should pay enough attention to the instruction of counting skills.

REFERENCES


Opdenakker, M.-C., & Van Damme, J. (2006). Differences between secondary schools: A study about school context, group composition, school practice, and school effects
with special attention to public and Catholic schools and types of schools. *School Effectiveness and School Improvement, 17*, 87-117.


