Cross-lingual activation in bilingual sentence processing: the role of word class meaning

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ABSTRACT

This study investigates how categorial (word class) semantics influences cross-lingual interactions when reading in L2. Previous homograph studies paid little attention to the possible influence of different word classes in the stimulus material on cross-lingual activation. The present study examines the word recognition performance of Dutch-English bilinguals who performed a lexical decision task to word targets appearing in a sentence. To determine the influence of word class meaning, the critical words either showed a word class overlap (e.g., the homograph tree [noun], which means ‘step’ in Dutch) or not (e.g., big [adj.], which is a noun in Dutch meaning ‘piglet’). In the condition of word class overlap, a facilitation effect was observed, suggesting that both languages were active. When there was no word class overlap, the facilitation effect disappeared. This result suggests that categorial meaning affects the word recognition process of bilinguals.
In the psycholinguistic research field dealing with bilingual lexical access, the dispute between selective and non-selective views seems to have been settled in favour of the latter. Despite early studies (Macnamara & Kushnir, 1971; Gerard & Scarborough, 1989) suggesting that access to the mental lexicon in bilingual word recognition is restricted to only one language, and that bilinguals are to a certain extent able to control their access to lexical information, other and more recent studies have shown that bilingual word recognition basically occurs automatically and in a non-selective way. This means that bilinguals, upon viewing a word, do not access the lexicon of each language separately but actually run a parallel search through their entire lexicon, irrespective of language. The most conclusive results stem from studies using cognates. In general, these studies unambiguously report that bilingual participants respond more quickly to L2 cognates than they do to L2 control words (e.g., Dijkstra, Grainger & Van Heuven, 1999; Lemhöfer & Dijkstra, 2004; Duyck et al., 2007). In contrast to the constant cognate facilitation effect, homograph studies yield a complicated pattern of results ranging from inhibition to no effects and to facilitation (Caramazza & Brones, 1979; Christoffanini, Kirsner & Milech, 1986; Beauvillain & Grainger, 1987; Dijkstra, Van Jaarsveld & Ten Brinke, 1998; Dijkstra, Grainger & van Heuven, 1999; Dijkstra, Timmermans & Schriefers, 2000; Dijkstra, De Bruijn, Schriefers & Ten Brinke, 2000; De Groot, Delmaar & Lupker, 2000; Lemhöfer & Dijkstra, 2004; Jared & Szucs, 2002; Von Studnitz & Green, 2002).

It is believed that the mixed and seemingly inconsistent results obtained in the homograph studies are due to experimental factors such as task demands and list composition. Yet, it is also argued that the inherent characteristics of the homographic items used could be responsible for the differential results. Dijkstra, Grainger and Van Heuven (1999), for instance, investigate the degree of cross-linguistic overlap on the basis of semantics (S),
orthography (O) and phonology (P), yielding six types: SOP, SO, SP, OP, O and P (see table 1).

[INSERT TABLE 1 ABOUT HERE]

The results show facilitation effects in the case of semantic and orthographic overlap (O and SO), while phonological overlap induces inhibition (P). For present purposes, only the homographic items O and OP are relevant. The researchers argue that the interplay of these two items might explain the previously observed null effects for homographs. The line of reasoning here is that the facilitatory effect due to the orthographic overlap is neutralised by the inhibitory effect due to the phonological overlap. Although a later study by Lemhöfer and Dijkstra (2004) did not report this phonology-ascribed inhibition, the findings in Dijkstra et al. (1999) suggest that null effects, and even inhibition effects, obtained in previous studies could be explained by the degree of cross-linguistic phonological overlap.

Previous studies thus provide different findings for cognates presented in isolation compared to homographs presented in isolation. But how are these words recognized in context? Does the presence of a sentence context, of the same language as the target word, restrict the activation of the non-target word, in the case of both cognates and homographs? Schwartz and Kroll (2006), for example, investigate the influence of a semantic context provided by a sentence on the recognition of cognates and homographs. To determine the influence of a semantic constraint on lexical activation, highly proficient and intermediate Spanish-English bilinguals had to name target and control words that were inserted in low-constraint and high-constraint sentences as quickly as possible (e.g. ‘we felt a bit nervous when we saw the fin of the shark in the distance’ and ‘from the beach we could see the shark’s fin pass through the water’, respectively). In low-constraint sentences, the results clearly demonstrate a cognate facilitation effect (replicated by Duyck et al., 2007 and Van Hell & De Groot, 2008). However, in the high-constraint sentences, recognition of
homographs and cognates did not produce a significant effect, nor did homographs in low-constraint sentences. Next to the fact that the results for homographs deviate again, the findings of Schwartz and Kroll (2006) establish two things: First, the cognate facilitation effect in low-constraint sentences shows that the presence of a unilingual sentence, does not restrict lexical search to a particular language. Second, the obtained null-effects in the other conditions demonstrate that the presence of a rich semantic context can eliminate lexical activation of the non-target language.

These findings are also established in studies using homophones. Li and Yip (1998), for example, examine cross-language homophones in the lexical processing of Chinese-English bilinguals. The homophones were either inserted in a predictive sentence context or in a neutral one. The results of their cross-modal naming experiment show that the predictive sentence context significantly facilitates Chinese-English bilinguals’ recognition of homophones, which in turn facilitates their naming of the phonologically related words. Similarly, in a recent eye-tracking study using English-French homophones (e.g. pool – poule), Chambers and Crooke (2009) demonstrate that a semantically compatible sentence context eliminates the activation of the English lexicon when interpreting French sentences. Once again, these studies provide evidence that the degree of non-selective activation is influenced by the semantic characteristics of the sentence context.

Comparable findings were also uncovered in a semantic priming paradigm. In a single-word semantic priming study, De Bruijn et al. (2001) report that regardless of the language, prime target words were responded to faster in a semantically related condition than in a semantically unrelated condition. This means that the language of context items in itself is not sufficient to suppress lexical candidates from the other language; only a semantic constraint is. Elston-Güttler (2000) adds that priming effects, found in single word lists, disappeared when homographs were presented at the end of full sentences, such as ‘The
woman gave her friend an expensive GIFT.’ This indicates that the semantic information provided by the preceding high-constraint sentence was sufficient to rule out the irrelevant meaning of the non-target language, German (Gift means ‘poison’). Later, Elston-Güttler et al. (2005) qualify this finding, arguing that bilinguals have to zoom into the all-L2 task. In their study, the results show a homograph priming effect in a sentence context, but only in the first half of the experiment. This is evidence that the participants gradually learn to deactivate the irrelevant meaning of the L1 reading of the homograph, based on the sentence context.

In sum, the studies on lexical access in a sentence context indicate that, in recognising interlingual homographs (or homophones for that matter), the presence of a sentence context could guide lexical access towards the target language. This especially applies to highly proficient bilinguals, whose results reveal an elimination of cross-lingual effects in a semantically high-constraint context (see also Elston-Güttler, Paulmann and Kotz, 2005). Semantic context here is to be defined on a general level. A recent lexical decision and translation study by Van Hell and De Groot (2008), however, examined the influence of semantics on cross-language activation at a more fine-grained level, by differentiating between concrete and abstract nouns. In this study, Dutch-English bilinguals had to perform a lexical decision task as well as translate target words (concrete and abstract cognates and noncognates) from L1 to L2 and vice versa. Targets were embedded in a high- or low-constraint context, or presented in isolation. The results reveal the same pattern as in Schwartz and Kroll (2006) and Duyck et al. (2007), being that the cognate facilitation effect only disappears in high-constraint sentences. It confirms that non-selectivity can be constrained by embedding words in a semantically constraining sentence context. In addition, the results show that this effect is sensitive to the more fine-grained variation of semantic overlap in backward and forward translation. Indeed, in the translation tasks, differences in concreteness modulate the interaction between cognate status and semantic constraint. This finding
indicates that it is worthwhile to go beyond the absolute all-or-none semantic overlap manipulations.

We believe that the degree of cross-linguistic overlap should be specified with respect to semantics, resulting in a clear differentiation between lexical semantics on the one hand, and word class semantics (or categorial meaning) on the other. Whereas Van Hell and De Groot’s (2008) study focuses on the former, the present manuscript will focus on the latter. We argue that the various results obtained in the homographs studies may be driven by the fact that part of the homograph stimuli share the same word class whereas others do not. The English words ANGEL and BREED, for example, are spelled identically to Dutch words meaning ‘sting’ and ‘wide’, respectively. However, ANGEL also shares the same word class in English and Dutch, being nouns in both languages, whereas BREED belongs to different word classes in the two languages. We assume that expectations regarding the word class will influence the degree of cross-lingual activation. This topic has so far received little empirical attention. A number of researchers acknowledge that little is known about the effects of semantic and syntactic information on cross-language activation (e.g., Dijkstra & van Heuven, 2002; French & Jacquet, 2004; Kroll & Tokowicz, 2005; van Hell, 2002).¹ Meuter (2009) actually mentions the lack of explicit claims concerning the influence of different word classes. In addition, Van Hell (2002) states that the effects of word class may constrain existing models of bilingual word recognition.² Recently, Sunderman and Kroll (2006) prove that it is relevant to control for word class. In a translation recognition study of words presented out of context, they show that bilinguals are sensitive to cues of word class, as form-related interference effects are reduced or eliminated when the word pairs are drawn from different categorial classes. The influence of word class meaning, however, becomes particularly relevant in a sentence context, as it is only in interaction with other words that meaning is fully deployed.

One may, for example, attribute a lexical-class meaning to the word fire, yet the categorial meaning of this word depends on actual speech and whether fire is used as a verb (for
example, in the sentence *They fired rubber bullets*) or as a noun (for example, in *There was a fire on the ground floor*) depends on the sentence as a whole. (Willems, 2000, p. 96)

Previous homograph recognition studies, presenting their items in a sentence context, did not control their stimuli for word class semantics, but only for semantics on a general level. Therefore, we systematically control for categorial semantics in this study. As far as we know, no other study has made this distinction with regard to the recognition of interlingual homographs in a sentence context.³

**Method**

**Participants**

We constructed a single experiment that was carried out by high-proficient Dutch-English bilinguals, being 7 male and 26 female university students ranging from 19 to 23 years old. This group received monetary compensation for participation. All participants were native speakers of Dutch living in an L1-dominant environment, speaking Dutch daily in their community. They have learnt English in a formal schooling setting from the age of 14 onwards and were, from an even younger age, more than regularly exposed to English (e.g., through the Internet, music, films, television and so forth). In addition to this daily exposure, participants were more exposed to English than is usual for the average speaker of Dutch, as they were all reading English language and literature at university. Of the 33 participants, 1 was excluded because of poor performance in the lexical decision task (his mean error rate was more than 2.5 standard deviations above the overall mean error rate).

**Stimulus Materials**

The critical target stimuli consisted of 32 English-Dutch homographs, of which 16 share their word class category over languages and 16 do not. Furthermore, the homographs were subdivided into a high-frequent and low-frequent group, based on their use in English (HFE
and LFE respectively). The homographs were extracted from the CELEX database (Baayen, Piepenbrock & Van Rijn, 1993). Using the WordGen stimulus generation program (Duyck et al., 2004), the control condition was created through an item-by-item match with respect to word frequency in English, neighbourhood size, comparable syllable structure (bigram) and word length (see table 2). Table 2 shows that the homographs and their control words did not differ from each other with regard to these variables. These tests prove that the English readings of the homographs and the controls only differ in terms of their interlingual homographic status.

[INSERT TABLE 2 ABOUT HERE]

An important criterion for the selection of critical words was that both the homograph and its control word had to be usable in the same low-constraint sentence as a final word (see table 3). Low constraint should be understood here on a lexical-semantic level. Nonetheless, the sentences do create expectation regarding the word class semantics of the targets. This is comparable to Schwartz and Kroll’s (2006) study in which the high-constraint environments create a certain expectation regarding the target’s lexical semantic field. In our study, expectancy is not based on “what a word means” (lexical semantics) but on “what word class a word belongs to” (categorial semantics).

[INSERT TABLE 3 ABOUT HERE]

As non-critical stimuli, we composed 16 low-constraint filler sentences with English filler target words and also 48 filler sentences with non-words as final words. These latter filler sentences were comparable to the sentences used for homographs and their respective controls. All 64 filler targets were again generated by using the WordGen stimulus generation program (Duyck et al., 2004), which produces non-words that are orthographically and phonologically conformable to English. In all, the stimulus list consisted of 96 experimental
trials: 16 homographs, 16 controls, 16 fillers and 48 non-words. This way, participants were
distracted from the presence of homographs. Furthermore, critical and non-critical sentences
were presented in random order so that the sequence of the sentences was completely
unpredictable.

After conducting the experiments, 3 critical pairs out of 32 were excluded from the
group’s RT-analyses. POMP (overlap), REIN and ROMP (both no overlap) and their controls
were discarded from further analysis because the mean error rate was more than 2.5 standard
deviations above the overall mean error rate.

**Procedure**

Participants were tested in small groups. They were placed in such a way that it was
impossible to see each other’s computer screens. Oral as well as written instructions were
given in English on how to perform an L2 lexical decision task to word targets appearing as
final words (serial visual presentation, SVP; see Schwartz & Kroll, 2006 and Duyck et al.,
2007). Participants were asked to respond by pressing one button if the word belonged to the
English language and another button if it did not. They were instructed to react as soon as
they identified the target as being an English word. It was ensured that participants were not
aware of the fact that L1 was crucial for the experiment; they were told that the experiment
was about L2 processing.

Our English sentences were presented in the middle of the computer screen, one word
at a time. Each word was presented in black print and centred for 700 ms on a flat-screen
monitor showing a white background. The upcoming appearance of the targets was indicated
by a preceding word highlighted in red, which also lingered on the middle of the screen (1200
ms). The red highlight was important in order to ensure word class disambiguation. One could
argue, for instance, that sentences with a noun target like “He trembled for a moment when he
saw the LOVER” cause a problem with regard to the expected word class because the target’s preceding article in this sample sentence could be followed by an adjective. However, the red print of the article makes such an expectation impossible. Targets were then presented in capital letters as a cue to respond with a maximum response time of 2500 ms. The interval between the successive presentations of sentences measured 1200 ms.

The presentation order of the 96 sentences was random and presented in 6 blocks of 10 sentences and 4 blocks of 9 sentences (preceded by 1 practice block of 10 sentences). Also, participants saw each sentence only once, either with the homograph or its respective control word as the target. To ensure that the participants actually read the sentences, we used a recognition task (similar to Elston-Güttler et al., 2005 and Duyck et al., 2007). After each block of 9 or 10 sentences, four sentences were presented, two of which were shown in the preceding block. Participants had to indicate for each of these four sentences whether it appeared in the preceding block. Mean accuracy on this verification task was very high (M = 85.59%, SD = 3.7).

Results

Statistical Methodology

A linear mixed model with fixed-effect terms for homograph (yes, no), categorial overlap (yes, no), frequency (LFE, HFE) and time (in which quarter of the experiment) and crossed random effects for subject and item (Baayen et al., 2008) was fitted to assess the cross-lingual activation of homographs. To explore the effect of word class overlap, the same mixed models were used but with an additional interaction term for overlap-by-homograph. Statistical analyses of the proportion of incorrect responses were carried out using a mixed-effects logistic regression (Jaeger, 2008) with the same terms as those used above. Analyses were carried out using R, an open-source programming language and environment for
statistical computing (R Development Core Team, 2007) and in particular the lme4 package for linear mixed-effects models (Bates, 2005).

**Reaction Time Analysis**

The single experiment constituted 32 participants and 29 items. The proportion of incorrect responses was 12.9%. These trials were excluded from all RT analyses. Also, RTs that were more than 2.5 standard deviations below or above the participant’s RT were excluded from the RT analysis (1.3% of the data).

[INSERT FIGURE 1 ABOUT HERE]

The bar plot in Figure 1 shows the mean RT in each condition of interest, averaged over individual trials, with the error bar representing the standard error of the mean. From these simple presentations, ignoring the dependency within subjects and words, trends in the data are quite clear but need to be confirmed by the statistical models accounting for these dependencies. Table 4 shows the results of the linear mixed-effect analyses.

[INSERT TABLE 4 ABOUT HERE]

A marginally significant effect of homographs vs. control on reaction time is observed (p=0.074). This difference is driven by words with categorial overlap (p=0.004). Reaction times for homographs with categorial overlap were 49 ms smaller compared to their controls, while no significant effect (p=0.706) was observed for homographs with no categorial overlap. These findings clearly demonstrate that the degree of cross-linguistic overlap in terms of word classes modulates cross-lingual activation effects.

Following Elston-Güttler et al. (2005), additional statistical analyses were carried out on the experimental data divided by half. The results show that the facilitation effect obtained for homographs with word class overlap was less pronounced in the second half (-35ms,
p=0.109) compared to the first half (-55ms, p=0.037). This might suggest that bilinguals zoom into the L2-task in that expectations regarding categorial meaning fade away. Nevertheless, the halving of the data makes statistical analyses less precise and reliable. Therefore, rather than interpreting the decreasing trend in RT as indicating a process of zooming into L2, it merely shows that participants learn how to perform on the specific experimental task.

**Error Rate Analysis**

The error rate (i.e., the number of incorrect responses divided by the total number of trial times), averaged over individual trials, is shown in Figure 2. The error rate for homographs is higher than is the error rate for controls, and it is driven by words with no categorial overlap.

[INSERT FIGURE 2 ABOUT HERE]

The results from the mixed-effects logistic model corroborate these findings. Effects in table 5 have to be interpreted on the log-odds ratio scale: the significant positive estimate of homograph versus the control effect, for example, indicates that this condition corresponds to a greater probability of an incorrect response. While the odds of an incorrect response (defined as the probability of an incorrect response divided by one minus this probability) is similar for the homographs and controls with categorial overlap, the odds are almost five times as high for the homographs with no overlap relative to their controls.

[INSERT TABLE 5 ABOUT HERE]

**Discussion**

Earlier research has shown that lexical access in bilingual homograph recognition, generally accepted as non-selective with respect to language, can be influenced by several factors, such as external experimental factors and stimuli-inherent differences in the degree of cross-linguistic overlap (i.e. SOP representations). Seeing that these findings are based on isolated
word recognition studies, Schwartz and Kroll (2006), among others, have examined the nature of bilingual lexical activation in a sentence context. They find that language membership of a sentence per se is not sufficient to constrain cross-lingual activity. Only a strong semantic context can suppress the activation of the non-target language. With semantics, Schwartz and Kroll (2006, p.209) refer to lexical-level and message-level semantic information. However, no distinction is made between lexical semantics and categorial semantics. Conceptually similar to the experiments conducted by Schwartz and Kroll, our study additionally examined whether the degree and direction of cross-lingual interactions might be affected by categorial semantics as well.

Overall, our findings demonstrated that a unilingual linguistic context is insufficient to constrain lexical activation to that specific language, as cross-lingual activation effects were obtained. The effects of non-selectivity in this experiment only decreased when the non-target could not tie in with the sentence—more specifically, with its expected categorial meaning—in the target language. Indeed, homograph facilitation effects were found in the case of word class overlap but were eliminated when there was no such overlap. As in Schwartz and Kroll’s (2006, p.208) study, “this suggests that the top-down processes of sentence comprehension can interact directly with the bottom-up process of lexical access and reduce the number of lexical entries that compete for selection.”

Our finding that categorial meaning modulates the extent of cross-lingual activations is compatible with the word recognition study by Dijkstra, Grainger and Van Heuven (1999), which investigates the degree of cross-linguistic overlap on the basis of lexical semantics (S), orthography (O) and phonology (P). Dijkstra et al. (1999) demonstrate that the usual facilitatory effect due to orthographic overlap is neutralised by the inhibitory effect due to phonological overlap. Analogously, we found that the lack of word class overlap has similar
inhibitory effects on homographic items. As in Sunderman and Kroll (2006) it shows that form-related effects are eliminated when the word class of the homographs is different.

Furthermore, our homograph facilitation effect is comparable to the cognate facilitation effect in low-constraint sentences (Schwartz & Kroll, 2006; Duyck et al., 2007; Van Hell & De Groot, 2008). Analogous to the convergence of orthography and lexical semantics, yielding facilitation effects in the cognate studies, we assume that the convergence of orthography and categorial semantics produced the facilitation effect in our study. In addition, it should be noted that Duyck et al. (2007) observe a stronger cognate facilitation effect with identical cognates compared to non-identical cognates. In other words, they show that the degree of cross-linguistic overlap indeed affects the strength of cross-lingual activations. In an unpublished doctoral thesis, Font (2001) demonstrates something alike, using cognates with orthographic difference at the end of the word (Fr. texte; Sp. texto) or in the middle (Fr. usuel; Sp. usual); the first group revealing a larger degree of facilitation than the latter. In these studies the degree of cross-linguistic overlap was defined in terms of orthography, while in ours, different categorial L1-L2 relationships were taken into account.

As for the accuracy data, our highly proficient participants made more errors on homograph items than on control items. Again, this effect is modulated by the degree of cross-lingual overlap—in exactly the opposite way of the RT analyses. Homographs with no overlap induced significantly more errors (typically on items associated with slower RT$^5$) compared to controls, while this was not the case with homographs with word class overlap. This shows that the activation of the non-targets occurs especially when the lexical competitors are only form-related.

Finally, it is important to discuss some implications of the present study for the future development of models of bilingual language processing. The most explicit model of visual word recognition in bilinguals is currently the BIA+ model of Dijkstra and Van Heuven
Following its predecessor (the BIA model, Dijkstra et al., 1999; Dijkstra & Van Heuven, 2002), BIA+ assumes an integrated L1/L2 lexicon with parallel bottom-up activation; only a minimal role is attributed to top-down influence. This means that non-linguistic context factors such as task demands and participant expectations can only indirectly affect the word identification system. Conversely, it is assumed that the presence of linguistic factors such as lexical, semantic and syntactic information offered by a sentence context could be able to directly influence the word recognition system, thus limiting non-selectivity. In other words, the modellers leave open the possibility that language membership of a target word is pre-activated by the presence of a sentence context and as a result will prevent the non-target from being activated. Schwartz and Kroll (2006), on the contrary, show that language membership exerts only minimal influence, as cross-lingual activation continues to occur. The present study replicated this finding and likewise provided support for the notion that language membership tags do not constrain non-selectivity when a sentence context is present. The available language nodes in BIA+ are therefore considered to be merely passive.

The major new finding of this study is that the effects of cross-lingual activation are influenced by whether or not language-ambiguous words share categorial meaning. Homographs belonging to the same word class over both languages are more rapidly responded to than those that do not. This finding is in harmony with the BIA+ model, as it assumes that the increased activation of semantics (through the presence of a sentence context which also indicates the categorial meaning) could potentially have a direct impact on cross-lingual activation. However, the BIA+ model does not fully specify which type of semantics is intended. Schwartz and Kroll (2006) already indicate this and propose a distinction between lexical-level and message-level semantics; Van Hell and De Groot (2008), for their part, offer a more fine-grained analysis on the lexical level. The present study provides evidence that
categorial semantics also plays a role. In addition to Van Hell and De Groot (2008), our study established that a broad notion such as semantics needs to be defined in greater detail in order to fully grasp the effect it has on (the degree of) cross-lingual activations.
Bibliography


In linguistics, there is some debate on whether to define word class as a semantic or a syntactic category. Since it is not the objective of this paper to discuss this issue, we merely assign word class to be a part of semantics.

2 See for instance Van Petten and Kutas (1991) who investigate lexical processing of open- versus closed-class words (i.e., content words such as verbs, nouns and adjectives versus function words such as prepositions and conjunctions).

3 Elston-Güttler and Friederici (2005) control for word class with noun-noun (same category; e.g. "fan") and noun-verb (mixed category; e.g. "trip") ambiguities, but they use intra-language homographs in English, testing English natives and German learners of English. Also Elston-Güttler et al. (2005) bring up word class semantics in passing, but they do not control for it. In fact, a closer look at their critical items showed that only 42 items reveal word class overlap, whereas the remaining 12 reveal word class contrasts.

4 A pilot experiment by Duyck et al. (2007) demonstrated that this was the rate at which participants (of similar level of our L2 proficiency participants) could comfortably process the L2 sentences.

5 We refer to Loeys, Rosseel and Baten (submitted) for more details on this joint modeling approach for reaction time and accuracy.
Table 1. Examples of word types used in Dijkstra, Grainger and van Heuven (1999).

<table>
<thead>
<tr>
<th>SOP</th>
<th>SO</th>
<th>SP</th>
<th>OP</th>
<th>O</th>
<th>P</th>
</tr>
</thead>
<tbody>
<tr>
<td>hotel</td>
<td>type</td>
<td>news</td>
<td>star</td>
<td>stage</td>
<td>cow</td>
</tr>
<tr>
<td>hotel</td>
<td>type</td>
<td>nieuws</td>
<td>star</td>
<td>stage</td>
<td>kou</td>
</tr>
</tbody>
</table>
Table 2. Mean lexical characteristics of stimuli used.

<table>
<thead>
<tr>
<th>Condition</th>
<th>Letters</th>
<th>Word frequency(^a)</th>
<th>Neighbourhood size(^b)</th>
<th>Bigram frequency(^c)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Homographs</td>
<td>4.13 (0.7)</td>
<td>1.29 (0.8)</td>
<td>8.22 (5.5)</td>
<td>6242.34 (3622.1)</td>
</tr>
<tr>
<td>Control words</td>
<td>4.13 (0.7)</td>
<td>1.34 (0.6)</td>
<td>8.31 (5.5)</td>
<td>5754.75 (3311.2)</td>
</tr>
<tr>
<td>p</td>
<td>identical</td>
<td>&gt;.72</td>
<td>&gt;.59</td>
<td>&gt;.32</td>
</tr>
</tbody>
</table>

Note. Standard deviations are indicated between brackets. Reported p values indicate significance levels of dependent samples comparisons between homographs and controls (matched item by item). L1 translations are displayed between square brackets.

\(^a\) Logarithm of word frequency per million words according to the CELEX lexical database (Baayen et al., 1993).

\(^b\) Neighbourhood size calculated using the WordGen program (Duyck et al., 2004).

\(^c\) Bigram frequency calculated using the WordGen program (Duyck et al., 2004).
Table 3. Examples of sentence stimuli used.

<table>
<thead>
<tr>
<th>Condition</th>
<th>Word class of homograph in English / Dutch</th>
<th>Target sentence (homograph / control word)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Overlap</td>
<td>N / N</td>
<td>She looked up and there seemed to be an ANGEL / ALIEN</td>
</tr>
<tr>
<td></td>
<td>N / N</td>
<td>Quickly and without thinking about it he pulled the LEVER / LEACH</td>
</tr>
<tr>
<td></td>
<td>N / N</td>
<td>You have to be careful because she is holding a WAND / HOSE</td>
</tr>
<tr>
<td></td>
<td>A / A</td>
<td>That guy can tell you with certainty that she is very GLAD / TIDY</td>
</tr>
<tr>
<td>No overlap</td>
<td>A / N</td>
<td>He told me he thinks this news is very BIG / SAD</td>
</tr>
<tr>
<td></td>
<td>V / N</td>
<td>All of a sudden those three kids started to JAM / HUM</td>
</tr>
<tr>
<td></td>
<td>A / N</td>
<td>My friend came in and said he would make it BRIEF / FANCY</td>
</tr>
<tr>
<td></td>
<td>N / A</td>
<td>She was amazed when she saw such type of BREED / STRAW</td>
</tr>
</tbody>
</table>